

# Relative Stability for Strictly Stationary Sequences

Zbigniew S. Szewczak

*Nicholas Copernicus University, Toruń, Poland*

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For a nonnegative strictly stationary random sequence satisfying the “minimal” dependence condition necessary and sufficient conditions for the relative stability are found. As an application the well-known Khinchine stability result for i.i.d. random variables is proved for uniformly strong mixing sequences. © 2001 Academic Press



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## 1. INTRODUCTION

Let  $\{X_k\}_{k \in \mathbb{Z}}$  be a strictly stationary sequence of nonnegative random variables defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Let  $S_n = \sum_{k=1}^n X_k$  and let  $v_n \rightarrow +\infty$  be a sequence of positive numbers.

For an independent identically distributed sequence of random variables Khinchine [12] defined the relative stability of  $\{X_k\}$  by the condition

$$c_n^{-1} S_n \rightarrow_P 1 \quad \text{as } n \rightarrow +\infty, \quad (1.1)$$

for some sequence  $c_n \rightarrow +\infty$ , and proved that Eq. (1.1) is equivalent to the slow variation of  $E(X_1 \wedge x)$  [3, Theorem 8.1.1; 6, Theorem 2, VIII, Sect. 9]. In such a case normalizing sequence  $c_n$  can be defined by

$$c_n \stackrel{\text{def}}{=} \inf \{x > 0; x^{-1} E(X_1 \wedge x) < 1/n\} \quad (1.2)$$

and is 1-regularly varying [3, Theorem 1.5.12, p. 28].

W. Feller [5, Chap. X, Sect. 4] uses Eq. (1.1) as a definition of “fair” game in the case of the absence of expectations and proves that for Petersburg paradox:

$$P(X_k = 2^i) = \frac{1}{2^i}, \quad i \geq 1,$$



one can take  $c_n = n \log_2 n$  (for the “fair” entrance fee problem see also [17]).

Raikov [16; 7, V, Sect. 28, Theorem 4] proved that the relative stability of  $\{X_k^2\}$  is equivalent to CLT. Since the Raikov Theorem remains valid for stationary martingale difference sequence (see [8, Theorem 3.2, p. 58]), the question of the relative stability (i.e., the existence of  $c_n \rightarrow +\infty$ , such that Eq. (1.1) holds) of an arbitrary nonnegative strictly stationary sequence is interesting. Of course, the case  $EX_1 < +\infty$  is covered by the Birkhoff Theorem so we assume throughout this paper that  $EX_1 = +\infty$ .

The method we use to prove relative stability results was introduced by Bernstein in [1] and is as follows

For a strictly stationary random sequence  $\{X_k\}$  let us define

$$\xi_j = \sum_{i=(j-1)(p+q)+1}^{jp+(j-1)q} X_i; \quad \eta_j = \sum_{i=jp+(j-1)q+1}^{j(p+q)} X_i$$

thus we have

$$v_n^{-1} S_n = v_n^{-1} \sum_{j=1}^k \xi_j + v_n^{-1} \sum_{j=1}^k \eta_j + v_n^{-1} \sum_{i=k(p+q)+1}^n X_i \quad (1.3)$$

where  $k(p+q) \sim n$ . If the last two summands on the r.h.s. of Eq. (1.3) tend in probability to 0 then the law  $\mathcal{L}(v_n^{-1} S_n)$  is asymptotically the same as the law  $\mathcal{L}(v_n^{-1} \sum_{j=1}^k \xi_j)$ . The latter is in turn asymptotically the same as the law  $\mathcal{L}^{*k}(v_n^{-1} \xi_1)$ -provided that  $\xi_j$  are *asymptotically independent*.

The above method, also known as a “big blocks by small blocks separation” requires the “dependence” condition  $B(v_n)$  [10],

$$\max_{1 \leq k+l \leq n} \left| E \left( \exp \left\{ it \frac{S_{k+l}}{v_n} \right\} \right) - E \left( \exp \left\{ it \frac{S_k}{v_n} \right\} \right) \cdot E \left( \exp \left\{ it \frac{S_l}{v_n} \right\} \right) \right| \rightarrow_n 0, \quad (1.4)$$

for some sequence  $v_n \rightarrow +\infty$  of nonnegative reals and every  $t \in \mathbb{R}$ . Since the measure degenerated at 1 is strictly stable thus by Theorem 3.1, [10], under relative stability, 1-regular variation of normalizing sequence  $v_n$  is equivalent to Eq. (1.4). In this sense Condition  $B$  is “minimal” and cannot be dropped.

Our criterion for relative stability will explore the notion of a sequence regularly varying in the limit.

Following [4] we shall say that the sequence of measurable, nonnegative functions  $\{f_n\}$  is  $(-\gamma)$ -regularly varying in the limit if there is a “rate”

sequence  $\{r_n\}$ ,  $r_n \rightarrow +\infty$ , such that for any sequence  $\{x_n\}$ , dominated by the rate sequence (i.e. such that  $x_n = o(r_n)$ ) and  $x_n \rightarrow +\infty$  we have

$$x_n^\gamma f_n(x_n) \rightarrow 1.$$

In the case when  $\gamma = 0$  we say that  $\{f_n\}$  is slowly varying in the limit.

A symmetric strictly stationary random sequence  $\{X_k\}$  satisfying condition  $B(v_n)$  is in the domain of attraction of symmetric strictly  $p$ -stable law,  $p \in (0, 2)$ , if, and only if

$$\left\{ \frac{x^p}{c} P(S_n > x v_n) \right\}$$

is slowly varying in the limit for some  $c > 0$  (Theorem 1, [4]). One can show ([19]) that for  $p = 2$  the corresponding condition is the slow variation in the limit of

$$\left\{ E \left( \frac{S_n^2}{v_n^2} \wedge x \right) \right\}.$$

The relative stability can be described in the same manner.

**THEOREM 1.** *Let  $\{X_k\}$  be a strictly stationary sequence of nonnegative random variables for which the Condition  $B(v_n)$  is fulfilled. Then the relative stability of  $\{X_k\}$  is equivalent to the slow variation in the limit of the sequence of functions*

$$\left\{ E \left( \frac{S_n}{v_n} \wedge x \right) \right\}.$$

While elegant Theorem 1 may be difficult to apply directly. It turns out that the version of Theorem 1 for “truncated” random variables is sometimes easier to verify. For a sequence  $\{c_n\}$ ,  $c_n \rightarrow +\infty$  let us define

$$X'_{n,k} = X_k I(X_k \leq c_n), \quad T'_{n,j} = \sum_{k=1}^j X'_{n,k}, \quad T'_n = T'_{n,n}, \quad \mathfrak{I}_n = E T'_n. \quad (1.5)$$

**THEOREM 2.** *Let  $\{X_k\}$  be a strictly stationary sequence of nonnegative random variables. Assume that one can find a sequence  $\{c_n\}$  of positive reals such that  $c_n \rightarrow +\infty$  and such that  $\mathfrak{I}_n$ , defined by Eq. (1.5) satisfies both Condition  $B(\mathfrak{I}_n)$  and*

$$\mathfrak{I}_n^{-1} \sum_{k=1}^n X_k I(X_k > c_n) \rightarrow_P 0 \quad \text{as } n \rightarrow +\infty. \quad (1.6)$$

Then

$$\mathfrak{g}_n^{-1} S_n \rightarrow_P 1, \quad \text{as } n \rightarrow +\infty, \quad (1.7)$$

if, and only if the sequence  $\{\mathfrak{g}_n^{-1} T'_n\}_n$  is uniformly integrable.

Let  $\mathcal{F}_k^m$  denote the  $\sigma$ -field generated by random variables  $\{X_i; k \leq i \leq m\}$  and let us define

$$\varphi_n = \varphi_n(\{X_k\}) = \sup \{|P(B | A) - P(B)|; A \in \mathcal{F}_{-\infty}^0, B \in \mathcal{F}_n^\infty\}.$$

$\{X_k\}$  is  $\varphi$ -mixing or uniformly strong mixing if  $\varphi_n \rightarrow 0$ . Clearly  $\varphi$ -mixing sequence satisfies Condition B with arbitrary normalizing sequence  $v_n \rightarrow +\infty$  (Proposition 5.2, [10]). We shall apply Theorem 1.6 to establish Khinchine stability result for  $\varphi$ -mixing sequences.

**THEOREM 3.** *Let  $\{X_k\}$  be a strictly stationary sequence of nonnegative  $\varphi$ -mixing random variables such that  $EX_1 = \infty$ .  $\{X_k\}$  is relatively stable if, and only if,  $E(X_1 \wedge x)$  is a slowly varying function in the sense of Karamata.*

Note that in the general situation we cannot avoid some additional condition for  $\{X_k\}$  (such as  $\varphi$ -mixing in Theorem 3).

**EXAMPLE 1.** Let  $X_k = X$  where  $X$  is a nonnegative random variable such that  $EXI(X \leq x)$  is a continuous slowly varying function and  $EX = +\infty$ . Since  $EXI(X \leq x) \sim E(X \wedge x)$  so  $c_n \sim \mathfrak{g}_n$  thus

$$\mathfrak{g}_n^{-1} S_n = \frac{X}{EXI(X \leq c_n)} \rightarrow_P 0,$$

and the sequence  $\{X_k\}$  satisfies Condition B( $\mathfrak{g}_n$ ). Observe that  $\{\mathfrak{g}_n^{-1} T'_n\}$  is not uniformly integrable because for any  $K > 0$

$$\mathfrak{g}_n^{-1} ET'_n I(T'_n > K\mathfrak{g}_n) = 1 - \frac{EXI(X \leq KEXI(X \leq c_n))}{EXI(X \leq c_n)} \rightarrow_n 1.$$

Of course, in Theorem 3 one cannot strengthen the convergence in probability to almost sure convergence.

**THEOREM 4.** *Let  $\{X_k\}$  be a strictly stationary sequence of nonnegative,  $\varphi$ -mixing random variables such that  $E(X_1) = \infty$  and  $E(X_1 \wedge x)$  is a slowly varying function. Let  $c_n$  be defined by Eq. (1.2). Then for any  $K > 0$*

$$\sum_{n=1}^{\infty} P(X_1 > Kc_n) = +\infty, \quad (1.8)$$

and

$$\limsup_n \frac{S_n}{c_n} = \limsup_n \max_{1 \leq k \leq n} \frac{X_k}{c_n} = \limsup_n \frac{X_n}{c_n} = +\infty \quad (1.9)$$

almost surely.

## 2. PROOFS

We begin with the following Proposition 1 parallel to Theorem 1 in [11].

**PROPOSITION 1.** *Let  $\{X_k\}$  be a strictly stationary sequence of non-negative random variables satisfying Eq. (1.4) with the sequence  $v_n \rightarrow +\infty$ . Then*

$$v_n^{-1} S_n \rightarrow_P 1 \quad \text{as } n \rightarrow +\infty,$$

*if and only if, there exists a sequence  $\{r_n\}$ ,  $r_n \rightarrow +\infty$  such that for every sequence  $x_n \rightarrow +\infty$ ,  $x_n = o(r_n)$  the following conditions are satisfied:*

$$x_n P(S_n > x_n v_n) \rightarrow 0 \quad \text{as } n \rightarrow +\infty, \quad (2.10)$$

$$x_n^{-1} \cdot v_n^{-2} \{ES_n^2 I(S_n \leq x_n v_n) - (ES_n I(S_n \leq x_n v_n))^2\} \rightarrow 0 \quad \text{as } n \rightarrow +\infty, \quad (2.11)$$

$$v_n^{-1} ES_n I(S_n \leq x_n v_n) \rightarrow 1 \quad \text{as } n \rightarrow +\infty. \quad (2.12)$$

*Proof of Proposition 1.* The proof mimics the proof of Theorem 1 in [11].

*Necessity.* For any  $n \in N$ , let  $\{Y_{n,i}\}_{i \in N}$  be a sequence of independent copies of  $v_n^{-1} S_n$ . By the hypothesis of Proposition 2.1 for each  $k \in N$

$$\mathcal{L} \left( k^{-1} \sum_{i=1}^k Y_{n,i} \right) \rightarrow_w \delta_1 \quad \text{as } n \rightarrow +\infty.$$

Therefore there exists a sequence of real numbers  $\{r_n\}$  such that  $r_n \rightarrow +\infty$ , and

$$\max_{1 \leq k \leq r_n} d_L \left( \mathcal{L} \left( k^{-1} \sum_{i=1}^k Y_{n,i} \right), \delta_1 \right) \rightarrow 0 \quad \text{as } n \rightarrow +\infty,$$

where  $d_L$  denotes the Lévy metric. Let  $\{k_n\}$  be a sequence of positive integers such that  $\lim k_n = \infty$ ,  $k_n = o(r_n)$ , then

$$\mathcal{L}\left(k_n^{-1} \sum_{i=1}^{k_n} Y_{n,i}\right) \rightarrow_w \delta_1 \quad \text{as } n \rightarrow +\infty,$$

where  $\{k_n^{-1} Y_{n,i}; 1 \leq i \leq k_n, n \in N\}$  is a sequence of infinitesimal, row-wise independent random variables. Equations (2.10), (2.11), (2.12) are satisfied by the Normal Convergence Criterion for  $\sigma^2 = 0$  (see [13, p. 328]).

The following lemma plays the same role as Lemma 1 in [11].

**LEMMA 1.** *Assume that there exists a sequence  $\{r_n\}$  of positive numbers with  $r_n \rightarrow +\infty$  as  $n \rightarrow +\infty$ , such that if  $x_n \rightarrow +\infty$  and  $x_n = o(r_n)$  then Eqs. (2.10)–(2.12) hold. If*

$$\mathcal{L}(\varsigma_n^{-1} S_n) \rightarrow_w \delta_1,$$

along subsequence  $n \in Q \subseteq N$  then

$$\varsigma_n v_n^{-1} \rightarrow_n 1$$

along  $n \in Q$ .

*Proof of Lemma 1.* There exists  $\{r'_n\}_{n \in Q}$ ,  $r'_n \rightarrow \infty$  such that for  $\{x'_n\}_{n \in Q}$ ,  $x'_n \rightarrow \infty$ ,  $x'_n = o(r'_n)$  we have

$$x'_n P(S_n > x'_n \varsigma_n) \rightarrow 0 \quad \text{and} \quad \varsigma_n^{-1} E S_n I(S_n \leq x'_n \varsigma_n) \rightarrow 1,$$

if  $n \rightarrow +\infty$ ,  $n \in Q$ . Suppose  $\lim x_n = +\infty$  and  $x_n = o(r_n \wedge r'_n)$  (where  $r_n$  is taken from Eqs. (2.10)–(2.12)) and  $\varsigma_n v_n^{-1} > 1 + \varepsilon$  along  $n \in Q' \subseteq Q$ . Integrating by parts we obtain

$$\begin{aligned} 1 &= \lim_{n \rightarrow \infty} \{x_n P(S_n > x_n \varsigma_n) + \varsigma_n^{-1} E S_n I(S_n \leq x_n \varsigma_n)\} \\ &= \lim_{n \rightarrow \infty} x_n \int_0^1 P(S_n > x_n \varsigma_n y) dy \\ &\leq (1 + \varepsilon)^{-1} \lim_{n \rightarrow \infty} ((1 + \varepsilon) x_n) \int_0^1 P(S_n > ((1 + \varepsilon) x_n) v_n y) dy \\ &= (1 + \varepsilon)^{-1} \lim_{n \rightarrow \infty} \{(1 + \varepsilon) x_n P(S_n > (1 + \varepsilon) x_n v_n) \\ &\quad + v_n^{-1} E S_n I(S_n \leq (1 + \varepsilon) x_n v_n)\} \\ &= (1 + \varepsilon)^{-1} \end{aligned}$$

because the sequence  $(1 + \varepsilon) x_n$  is dominated by the rate sequence  $r_n \wedge r'_n$ . Hence we have

$$\limsup_{n \in Q'} \frac{\zeta_n}{v_n} \leq 1.$$

Since  $\zeta$  and  $v$  can be exchanged in the above, Lemma 2.4 holds.  $\blacksquare$

*Sufficiency.* Let  $\{Y_{n,j}, j \in N, n \in N\}$  be a sequence of row-wise i.i.d. random variables such that  $\mathcal{L}(Y_{n,j}) = \mathcal{L}(v_n^{-1} S_n)$ . By the Normal Convergence Criterion for  $\sigma^2 = 0$  (see [13], p. 328) Eqs. (2.10), (2.11) and (2.12) imply

$$\mathcal{L}\left(k_n^{-1} \sum_{j=1}^{k_n} Y_{n,j}\right) \rightarrow_w \delta_1, \quad \text{as } n \rightarrow +\infty \quad (2.13)$$

for any sequence  $\{k_n\}_{n \in N}$  of positive integers such that  $\lim k_n = +\infty$  and  $k_n = o(r_n)$ .

The estimate, valid for fixed  $k \in N$ ;

$$\begin{aligned} & \left| E\left(\exp\left\{itv_{kn}^{-1} \sum_{j=1}^{kn} X_j\right\}\right) - \left(E\left(\exp\left\{itv_{kn}^{-1} S_n\right\}\right)\right)^k \right| \\ &= \left| E\left(\exp\left\{itv_{nk}^{-1} \sum_{j=1}^k \left(\sum_{i=1}^n X_{n(j-1)+i}\right)\right\}\right) - \prod_{j=1}^k E\left(\exp\left\{itv_{kn}^{-1} S_n\right\}\right) \right| \\ &\leq \sum_{l=1}^{k-1} \left| E\left(\exp\left\{itv_{nk}^{-1} \sum_{j=1}^{k-l+1} \left(\sum_{i=1}^n X_{n(j-1)+i}\right)\right\}\right) \cdot \prod_{j=1}^{l-1} E\left(\exp\left\{itv_{nk}^{-1} S_n\right\}\right) \right. \\ &\quad \left. - E\left(\exp\left\{itv_{kn}^{-1} \sum_{j=1}^{k-l} \left(\sum_{i=1}^n X_{n(j-1)+i}\right)\right\}\right) \cdot \prod_{j=1}^l E\left(\exp\left\{itv_{kn}^{-1} S_n\right\}\right) \right| \\ &\leq \sum_{l=1}^{k-1} \left| E\left(\exp\left\{it \frac{S_{(k-l+1)n}}{v_{nk}}\right\}\right) - E\left(\exp\left\{it \frac{S_{(k-l)n}}{v_{kn}}\right\}\right) \cdot E\left(\exp\left\{it \frac{S_n}{v_{kn}}\right\}\right) \right| \\ &\leq k \max_{1 \leq p+q \leq kn} \left| E\left(\exp\left\{it \frac{S_{p+q}}{v_{kn}}\right\}\right) - E\left(\exp\left\{it \frac{S_p}{v_{kn}}\right\}\right) \cdot E\left(\exp\left\{it \frac{S_q}{v_{kn}}\right\}\right) \right| \end{aligned}$$

and Eq. (1.4) yield

$$\mathcal{L}\left(\frac{S_{k_n n}}{v_{k_n n}}\right) \sim \mathcal{L}^{*k_n}\left(\frac{S_n}{v_{k_n n}}\right), \quad (2.14)$$

for  $k_n \rightarrow +\infty$  and new “rate” sequence  $\{r'_n\}$ .

We shall prove that Eq. (2.10) gives tightness of  $\{v_{k_n n}^{-1} S_{k_n n}\}$  and Eq. (2.12) implies that any weak limit is not degenerated at 0. To see this assume first that  $\{v_{k_n n}^{-1} S_{k_n n}\}$  is not tight. Then we have

$$\lim_{k \rightarrow \infty} \limsup_{n \rightarrow \infty} P(v_n^{-1} S_n > k) \geq \lim_{k \rightarrow \infty} \limsup_{n \rightarrow \infty} P(v_{k_n n}^{-1} S_{k_n n} > k) \geq \eta > 0.$$

Let  $N_l > N_{l-1}$  be a sequence of positive integers such that

$$P(v_{N_l}^{-1} S_{N_l} > l) > \frac{1}{2} \eta$$

and the index  $l$  is such that  $r_n > l^2$  for  $n \geq N_l$ . Define  $k_n = l$  for  $N_l \leq n < N_{l+1}$ . Observe that  $\{k_n\}$  satisfies

$$\frac{k_n}{r_n} \leq \frac{1}{k_n} \rightarrow 0$$

and

$$\liminf_{l \rightarrow \infty} k_{N_l} P(v_{N_l}^{-1} S_{N_l} > k_{N_l}) \geq \liminf_{l \rightarrow \infty} k_{N_l} \cdot \frac{1}{2} \eta = +\infty$$

which contradicts Eq. (2.10) and thus proves the tightness of  $\{v_n^{-1} S_n\}$  (and  $\{v_{k_n n}^{-1} S_{k_n n}\}$ ). To verify that each weak limit of  $\{v_{k_n n}^{-1} S_{k_n n}\}$  is not degenerated at 0 it is enough to verify the same for the sequence  $\{v_n^{-1} S_n\}$ . Assume that this is not the case, i.e. that along some subsequence  $n'$  we have

$$\mathcal{L}(v_{n'}^{-1} S_{n'}) \rightarrow_w \delta_0.$$

Then there exists a sequence  $k_{n'} = o(r_{n'})$  such that

$$\lim_{n' \rightarrow \infty} v_{n'}^{-1} E(S_{n'} I(S_{n'} \leq k_{n'} v_{n'})) = 0$$

which contradicts Eq. (2.12) and proves that any weak limit of  $\{v_n^{-1} S_n\}$  is not degenerated at 0.

Assume that  $Z$  is a limit along the subsequence  $\{k_{n'}\}$  such that  $\mathcal{L}(Z) \neq \delta_0$  and

$$\mathcal{L}\left(\frac{S_{k_{n'} n'}}{v_{k_{n'} n'}}\right) \rightarrow_w \mathcal{L}(Z). \quad (2.15)$$

By Eq. (2.14) we have

$$\mathcal{L}\left(\frac{v_{n'}}{v_{k_{n'} n'}} \sum_{j=1}^{k_{n'}} Y_{n', j}\right) \rightarrow_w \mathcal{L}(Z). \quad (2.16)$$



Suppose that along subsubsequence  $\{n''\} \subset \{n'\}$  we have

$$\frac{k_{n''} v_{n''}}{v_{k_{n''} n''}} \rightarrow_{n''} c.$$

Now if  $c = 0$  then by Eqs. (2.16), (2.13) we have that  $\mathcal{L}(Z) = \delta_0$ . If  $c = +\infty$  then Eqs. (2.13), (2.16), (2.14) imply that the sequence  $\{v_{k_n n}^{-1} S_{k_n n}\}$  is not tight—which is not possible as we have proved above. So we may assume that  $0 < c < +\infty$  and

$$0 < c_1 = \liminf_{n'} \frac{k_{n'} v_{n'}}{v_{k_{n'} n'}} < \limsup_{n'} \frac{k_{n'} v_{n'}}{v_{k_{n'} n'}} = c_2 < +\infty.$$

By this and Eqs. (2.16), (2.13) we have for  $0 < c_1 < c_2 < +\infty$

$$\mathcal{L}(c_1^{-1} Z) = \mathcal{L}(c_2^{-1} Z),$$

hence for  $c = \frac{c_1}{c_2} < 1$

$$\mathcal{L}(Z) = \mathcal{L}(cZ) = \mathcal{L}(c^2 Z) = \dots = \dots = \lim_{n \rightarrow \infty} \mathcal{L}(c^n Z) = \delta_0.$$

The above again is not possible since we have proved that  $\mathcal{L}(Z) \neq \delta_0$  so finally  $\mathcal{L}(Z) = \delta_c$  for some  $0 < c < +\infty$ . By Eq. (2.15) let  $r''_n \rightarrow +\infty$  be such that

$$\mathcal{L}\left(l_{n'}^{-1} \sum_{j=1}^{l_{n'}} Y_{n'k_n, j}\right) \rightarrow_w \mathcal{L}(Z),$$

for every  $l_{n'} \rightarrow +\infty$ ,  $l_{n'} = o(r''_{n'})$ . In particular, if  $l_{n'} = o(r''_{n'} \wedge r_{n'k_n})$ , Eqs. (2.10)–(2.12) imply  $\mathcal{L}(Z) = \delta_1$ . This proves that  $c = 1$  and so

$$v_{n \cdot k_n}^{-1} v_n k_n \rightarrow 1 \quad \text{as } n \rightarrow +\infty.$$

This, Eqs. (2.13) and (2.14) imply

$$\mathcal{L}((v_{nk_n})^{-1} S_{nk_n}) \rightarrow_w \delta_1 \quad \text{as } n \rightarrow +\infty, \quad (2.17)$$

for any  $k_n \rightarrow \infty$ ,  $k_n = o(r_n \wedge r'_n)$ .

In order to finish the proof we need to show that one can approximate  $S_n$  by  $S_{[(n+k_n)k_n^{-1}]k_n}$ . Let us define  $a_n = \min_{i \geq n} v_i$  and choose the sequence  $\{m_n\}$  of positive integers such that,  $\lim m_n = \infty$  and

$$\max_{1 \leq j \leq m_n} \frac{S_j}{a_n} \rightarrow_P 0 \quad \text{as } n \rightarrow +\infty.$$

For the rate sequence  $r_n \wedge r'_n \wedge m_n$  we have

$$\begin{aligned} & P(v_{[(n+k_n)k_n^{-1}]k_n}^{-1}(S_{[(n+k_n)k_n^{-1}]k_n} - S_n) > \varepsilon) \\ & \leq P(\max_{1 \leq j \leq k_n} a_n^{-1} S_j > \varepsilon) \rightarrow 0 \end{aligned}$$

as  $n \rightarrow +\infty$ . By Lemma 1 we have to prove that

$$\mathcal{L}\left(v_{[(n+k_n)k_n^{-1}]k_n}^{-1} S_{[(n+k_n)k_n^{-1}]k_n}\right) \rightarrow_w \delta_1 \quad \text{as } n \rightarrow +\infty.$$

In order to use Eq. (2.17) we have to find  $k_n$  and  $s_n$  such that

$$\frac{k_n}{S_{[(n+k_n)k_n^{-1}]}} \rightarrow 0 \quad \text{as } n \rightarrow +\infty.$$

Define the rate  $s_n = r_n \wedge r'_n \wedge m_n \wedge \sqrt{n}$  and observe that  $s_n \nearrow +\infty$  ( $r_n, r'_n, m_n$  can be chosen increasing),  $ns_n^{-1} \rightarrow +\infty$  and if  $k_n = o(s_{[ns_n^{-1}]})$  and  $k_n \rightarrow +\infty$  then,

$$\frac{k_n}{S_{[(n+k_n)k_n^{-1}]}} \leq \frac{k_n}{S_{[nk_n^{-1}]}} \leq \frac{k_n}{S_{[ns_n^{-1}]}} \rightarrow 0 \quad \text{as } n \rightarrow +\infty.$$

This completes the proof of Proposition 1. ■

### *Proof of Theorem 1.*

*Necessity.* Since by Lebesgue dominated convergence theorem  $E(\frac{S_n}{v_n} \wedge m) \rightarrow 1$  for every  $m > 1$  one can apply Corollary 8.5, [10].

*Sufficiency.* Assume that for some sequence  $\{x_n\}$ ,  $x_n \rightarrow +\infty$

$$\int_0^{x_n} P(S_n > yv_n) dy = E\left(\frac{S_n}{v_n} \wedge x_n\right) \rightarrow 1.$$

Let  $y_n = o(x_n)$ ,  $y_n \rightarrow \infty$ , then

$$\int_{y_n}^{x_n} P(S_n > yv_n) dy = \left\{ E\left(\frac{S_n}{v_n} \wedge x_n\right) - E\left(\frac{S_n}{v_n} \wedge y_n\right) \right\} \rightarrow_n 0.$$

By this we find that

$$x_n \left(1 - \frac{y_n}{x_n}\right) P(S_n > x_nv_n) \leq \int_{y_n}^{x_n} P(S_n > yv_n) dy \rightarrow 0,$$

hence

$$x_n P(S_n > x_n v_n) \rightarrow 0, \quad (2.18)$$

so Eq. (2.10) holds. On the other hand by the following identity

$$E \frac{S_n}{v_n} I\left(\frac{S_n}{v_n} \leq x_n\right) = -x_n P(S_n > x_n v_n) + \int_0^{x_n} P(S_n > y v_n) dy$$

for  $x_n = o(r_n)$ ,  $x_n \rightarrow \infty$  we get

$$E\left(\frac{S_n}{v_n} I\left(\frac{S_n}{v_n} \leq x_n\right)\right) \rightarrow 1,$$

so Eq. (2.12) holds. Thus it remains to establish Eq. (2.11). Since

$$E\left(\frac{S_n}{v_n} I\left(\frac{S_n}{v_n} \leq x_n\right)\right) \rightarrow 1,$$

thus

$$x_n^{-1} \left( E\left(\frac{S_n}{v_n} I\left(\frac{S_n}{v_n} \leq x_n\right)\right) \right)^2 \rightarrow 0,$$

so we need to show that

$$x_n^{-1} E\left(\frac{S_n^2}{v_n^2} I\left(\frac{S_n}{v_n} \leq x_n\right)\right) \rightarrow 0.$$

For this observe that

$$E\left(\frac{S_n^2}{v_n^2} I\left(\frac{S_n}{v_n} \leq x_n\right)\right) = -x_n^2 P(S_n > x_n v_n) + 2 \int_0^{x_n} y P(S_n > y v_n) dy,$$

so by Eq. (2.18) it suffices to show that

$$x_n^{-1} \int_0^{x_n} y P(S_n > y v_n) dy \rightarrow 0. \quad (2.19)$$

We have

$$\begin{aligned} & \int_0^{x_n} y P(S_n > y v_n) dy \\ &= \int_0^{x_n} y dE\left(\frac{S_n}{v_n} \wedge y\right) = x_n E\left(\frac{S_n}{v_n} \wedge x_n\right) - \int_0^{x_n} E\left(\frac{S_n}{v_n} \wedge y\right) dy, \end{aligned}$$

by integration by parts. Now by Theorem 8.2, [10] and assumptions of the theorem we obtain

$$x_n^{-1} \int_0^{x_n} E \left( \frac{S_n}{v_n} \wedge y \right) dy \rightarrow 1,$$

so Eq. (2.19) holds. Thus from Eq. (2.19) follows Eq. (2.11) and Theorem 1 is proved. ■

*Proof of Theorem 2.*

*Necessity.* Eqs. (1.6) and (1.7) imply  $\vartheta_n^{-1} T'_n \rightarrow_P 1$ . Uniform integrability of  $\{\vartheta_n^{-1} T'_n\}$  follows from Theorem 5.4, [2], p. 32.

*Sufficiency.* Let  $\{\vartheta_n^{-1} T'_n\}$  be uniformly integrable, so by Lemma 1, [11], we obtain for  $x_n \rightarrow \infty$ ,  $x_n = o(r_n)$

$$E \frac{T'_n}{\vartheta_n} I \left( \frac{T'_n}{\vartheta_n} > x_n \right) = x_n P(T'_n > x_n \vartheta_n) + \int_{x_n}^{+\infty} P(T'_n > y \vartheta_n) dy \rightarrow 0,$$

which proves that

$$x_n P(T'_n > x_n \vartheta_n) \rightarrow_n 0.$$

By this and the identity

$$E \frac{T'_n}{\vartheta_n} I \left( \frac{T'_n}{\vartheta_n} \leq x_n \right) = -x_n P(T'_n > x_n \vartheta_n) + \int_0^{x_n} P(T'_n > y \vartheta_n) dy \rightarrow 1,$$

we have

$$\int_0^{x_n} P(T'_n > y \vartheta_n) dy \rightarrow 1.$$

Since for every  $k \in N$  and  $\varepsilon > 0$  by Eq.(1.6)

$$\int_0^k P(S_n - T'_n > \varepsilon y \vartheta_n) dy \rightarrow 0,$$

thus there exists “rate”  $r'_n$ ,  $r'_n \leq r_n$  such that for  $x_n = o(r'_n)$

$$\begin{aligned} 0 &\leq \int_0^{x_n} P(S_n > y \vartheta_n) dy - \int_0^{x_n} P(T'_n > y \vartheta_n) dy \\ &\leq \int_0^{x_n} P(S_n - T'_n > \varepsilon y \vartheta_n) dy + \frac{1}{1-\varepsilon} \int_0^{(1-\varepsilon)x_n} P(T'_n > y \vartheta_n) dy \\ &\quad - \int_0^{x_n} P(T'_n > y \vartheta_n) dy \rightarrow \frac{1}{1-\varepsilon} - 1. \end{aligned}$$

Hence we proved that

$$\left\{ E \left( \frac{S_n}{g_n} \wedge x \right) \right\} = \left\{ \int_0^x P(S_n > y g_n) dy \right\}$$

is slowly varying in the limit and by Theorem 1 Eq. (1.7) holds. ■

*Proof of Theorem 3.*

*Necessity.* Assume that  $v_n^{-1} S_n \rightarrow_P 1$ . We shall prove first that

$$nP(X_1 > 2v_n) \rightarrow_n 0. \quad (2.20)$$

Let  $p$  be such that  $\varphi_p < 1$ . By Proposition 3.1, [15] for  $n > p$  we have

$$P(\max_{1 \leq k \leq [n/p]} X_{kp} > x) \geq (1 - \varphi_p) P(\max_{1 \leq k \leq [n/p]} X_{kp}^* > x),$$

where  $\{X_k^*\}$  is an i.i.d. sequence such that  $\mathcal{L}(X_1) = \mathcal{L}(X_k^*)$ . Hence for  $n > p$

$$P(\max_{1 \leq k \leq n} X_k \leq x) \leq 1 - (1 - \varphi_p) P(\max_{1 \leq k \leq [n/p]} X_{kp}^* > x).$$

This and nonnegativity of  $X_k$  and Eq. (1.1) imply

$$\begin{aligned} 1 &= \liminf_n P(S_n \leq 2v_n) \leq \liminf_n P(\max_{1 \leq k \leq n} X_k \leq 2v_n) \\ &\leq \liminf_n \{1 - (1 - \varphi_p) P(\max_{1 \leq k \leq [n/p]} X_k^* > 2v_n)\} \\ &\leq \varphi_p + (1 - \varphi_p) \liminf_n P^{[n/p]}(X_1 \leq 2v_n). \end{aligned}$$

Thus by the elementary inequality

$$x^n \leq \exp\{-n(1-x)\}, \quad 0 < x \leq 1,$$

we have that

$$1 - \varphi_p \leq (1 - \varphi_p) \exp \left\{ - \limsup_n \left[ \frac{n}{p} \right] P(X_1 > 2v_n) \right\},$$

hence Eq. (2.20) holds.  $v_n^{-1} S_n \rightarrow_P 1$  and  $v_n \rightarrow_n +\infty$  imply

$$\frac{v_{n+1}}{v_n} \rightarrow_n 1. \quad (2.21)$$

By Eqs. (2.20) and (1.1) with  $c_n = v_n$  we have for  $0 < \varepsilon < 1$

$$1 = \liminf_n P\left(\sum_{k=1}^n X_k I(X_k \leq 2v_n) > (1 - \varepsilon)v_n\right) \leq \liminf_n \frac{nEX_1 I(X_1 \leq 2v_n)}{(1 - \varepsilon)v_n}.$$

Hence

$$\frac{2v_n P(X_1 > 2v_n)}{EX_1 I(X_1 \leq 2v_n)} = \frac{2nP(X_1 > 2v_n)}{\frac{n}{v_n} EX_1 I(X_1 \leq 2v_n)} \rightarrow_n 0.$$

Thus Eq. (2.21) implies that for  $2v_n \leq x < 2v_{n+1}$

$$\frac{xP(X_1 > x)}{EX_1 I(X_1 \leq x)} \leq \frac{2v_{n+1}P(X_1 > 2v_n)}{EX_1 I(X_1 \leq 2v_n)} \rightarrow_n 0,$$

which proves the slow variation of  $EX_1 I(X_1 \leq x)$  and  $E(X_1 \wedge x)$  [6, Theorem 2, VIII, Sect. 9]. Observe that in the proof we used the condition  $\lim_n \varphi_n < 1$  only.

*Sufficiency.* Let  $c_n$  be defined as in Eq. (1.2). We have for any  $\varepsilon > 0$

$$\begin{aligned} P\left(\mathfrak{g}_n^{-1} \sum_{k=1}^n X_k I(X_k > c_n) > \varepsilon\right) &\leq nP(X_1 > c_n) \\ &= \frac{c_n P(X_1 > c_n)}{c_n P(X_1 > c_n) + EX_1 I(X_1 \leq c_n)} \cdot \frac{n}{c_n} E(X_1 \wedge c_n) \rightarrow_n 0, \end{aligned}$$

by Theorem 2, VIII, Sect. 9 of [6] and the definition of  $c_n$  so Eq. (1.6) holds. According to Theorem 2 it is enough to verify uniform integrability of  $\{\mathfrak{g}_n^{-1} T'_n\}$ . Observe that

$$\frac{xP(X_1 > x)}{E(X_1 \wedge x)} \leq \frac{2(E(X_1 \wedge x) - E(X_1 \wedge (2^{-1}x)))}{E(X_1 \wedge x)} \rightarrow 0,$$

as  $x \rightarrow +\infty$  so by

$$E(X_1 \wedge x) = xP(X_1 > x) + EX_1 I(X_1 \leq x),$$

we get  $EX_1 I(X_1 \leq x) \sim E(X_1 \wedge x)$  and  $\mathfrak{g}_n \sim c_n$ . Thus the sequence

$$\left\{ \max_{1 \leq i \leq n} \mathfrak{g}_n^{-1} X_i I(X_i \leq c_n) \right\}_n$$

is bounded a.s. and so is uniformly integrable. By the Markov inequality and stationarity we have

$$\max_{1 \leq i \leq n} P((T'_n - T'_{n,i}) > \vartheta_n 2^{-1} b x) \leq \max_{1 \leq i \leq n} (\vartheta_n b x)^{-1} 4E(T'_{n,i}) = \frac{4}{bx}.$$

By  $\varphi$ -mixing and the fact that r.h.s. of the above can be arbitrary small for large  $x$  and  $b$  we obtain that there exist  $a_0 > 0$ ,  $b > 0$ ,  $p \in N$  and  $\eta < \frac{1}{2}$ , such that the condition

$$\varphi_p + \max_{1 \leq i \leq n} P(|T'_n - T'_{n,i}| \geq \vartheta_n 2^{-1} b a_0) \leq \eta < 1.$$

is fulfilled together with

$$\frac{\eta(1+2b)}{(1-\eta)} < 1. \quad (2.22)$$

By Lemma 3.1, [14] we get for  $n > p$  i  $x > a_0$

$$\begin{aligned} P(T'_n > (1+2b) x \vartheta_n) &\leq (1-\eta)^{-1} P\left(\max_{1 \leq i \leq n} X'_{n,i} > (2p)^{-1} b x \vartheta_n\right) \\ &\quad + \eta \cdot (1-\eta)^{-1} P(T'_n > x \vartheta_n). \end{aligned}$$

Integrating both sides and using well known formula

$$EXI(X > x) = xP(X > x) + \int_x^{+\infty} P(X > y) dy,$$

we conclude that for  $K > a_0$

$$\begin{aligned} E \frac{T'_n}{\vartheta_n} I\left(\frac{T'_n}{\vartheta_n} > (1+2b) K\right) &\leq \frac{\eta(1+2b)}{(1-\eta)} E \frac{T'_n}{\vartheta_n} I\left(\frac{T'_n}{\vartheta_n} > K\right) \\ &\quad + \frac{2p(1+2b)}{b(1-\eta)} E \max_{1 \leq i \leq n} \frac{X'_{n,i}}{\vartheta_n} I\left(\max_{1 \leq i \leq n} \frac{X'_{n,i}}{\vartheta_n} > (2p)^{-1} b K\right). \end{aligned}$$

Taking  $\sup_n$  over both sides and using the fact that

$$\sup_n E \frac{T'_n}{\vartheta_n} I\left(\frac{T'_n}{\vartheta_n} > K\right)$$

is decreasing function of  $K$  and going with  $K \rightarrow +\infty$  we get from Eq. (2.22) that  $\vartheta_n^{-1} T'_n$  is uniformly integrable. ■

*Proof of Theorem 4.* We shall establish firstly that, if  $EX_1 = +\infty$  and  $E(X_1 \wedge x)$  is 1-regularly varying function then, the condition (1.8) is satisfied. Assume that this is not true, i.e., for some  $K > 0$

$$\sum_{n=1}^{\infty} P(X_1 > Kc_n) < +\infty.$$

Then by the fact that  $\{c_n\}$  is 1-regularly varying and by Lemma 3.2.4 ([18], p. 131), we obtain

$$\sum_{n=1}^{\infty} P(X_1 > Kc_n) < +\infty,$$

for every  $K > 0$ . By Feller's Theorem (see Theorem 3.2.5, [18], p. 132) for independent random variables  $\{X_k^*\}$  such that,  $\mathcal{L}(X_k^*) = \mathcal{L}(X_1)$  we have

$$\limsup_n \frac{\sum_{k=1}^n X_k^*}{c_n} = 0,$$

a.s. This contradicts Eq. (1.1) which holds by Khinchine's Theorem. Thus for every  $K > 0$

$$\sum_{n=1}^{\infty} P(X_1 > Kc_n) = +\infty.$$

By Eq. (1.8) the converse Borel–Cantelli lemma for  $\varphi$ -mixing sequences (see Lemma 1.1.2', [9], p. 3) can be applied. Thus we have for  $\varphi_p$  and  $K > 0$

$$P\left(\bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} \{X_n > Kc_n\}\right) \geq P\left(\bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} \{X_{np} > Kc_{np}\}\right) \geq 1 - \varphi_p,$$

which by the choice of  $K$  and for  $\varphi_p \rightarrow_p 0$  proves the last equality in Eq. (1.9) while the nonnegativity of  $X_k$  yields the rest. ■

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